

LOCAL FLUCTUATIONS OF CRITICAL MANDELBROT CASCADES

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ABSTRACT. We investigate so-called generalized Mandelbrot cascades at the freezing (critical) temperature. It is known that, after a proper rescaling, a sequence of multiplicative cascades converges weakly to some continuous random measure. Our main question is how the limiting measure μ fluctuates. For any given point x , denoting by $B_n(x)$ the ball of radius 2^{-n} centered around x , we present optimal lower and upper estimates of $\mu(B_n(x))$ as $n \rightarrow \infty$.

1. INTRODUCTION

1.1. Mandelbrot cascades. ¹ In the seventies Mandelbrot [31, 32] proposed a model of random multiplicative cascade measures, to simulate the energy dissipation in intermittent turbulence. Mandelbrot cascades exhibited a number of fractal and statistical features observed experimentally in a turbulence flow. Up to now, through various applications, this model found its way into a wide range of scientific fields from financial mathematics [15] to quantum gravity and disordered systems in mathematical physics [3]. Mathematically, a multiplicative cascade, is a measure-valued stochastic process and was first rigorously described by Kahane and Peyrière [27]. They presented a complete proof of results announced by Mandelbrot, answering e.g. the questions of non-degeneracy, existence of moments and local properties. Since then multiplicative cascades become a subject of study for numerous mathematicians, see e.g. [6, 7, 17, 23, 29].

One of the simplest examples of multiplicative cascades can be expressed as a sequence of random measures on the unit interval $I = [0, 1)$. They depend on two parameters: a real number $\beta > 0$ (inverse temperature parameter) and a real valued random variable ξ (fluctuations). For convenience, we assume that ξ is normalized, i.e.

$$(1.1) \quad \mathbb{E}e^\xi = \frac{1}{2} \quad \text{and} \quad \mathbb{E}[\xi e^\xi] = 0.$$

To define the cascade measures, consider an infinite dyadic Ulam-Harris tree denoted by $\mathbb{T}_2 = \bigcup_{n \geq 0} \{0, 1\}^n$ and attach to every edge, connecting x with xk ($x \in \mathbb{T}_2$, $k \in \{0, 1\}$), a random weight $\xi_k(x)$, being an independent copy of ξ . Let $V(x)$ be the total weight of the branch from the root to x obtained by adding weights of the edges along this path.

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¹In order to keep the introduction as accessible as possible, we omit some technical subtleties. This way we focus on the general overview and the contribution of this paper. We postpone the proper introduction of the setting and notation to the second section.

Define the measure $\mu_{\beta,n}$ on the unit interval I as an absolutely continuous with respect to the Lebesgue measure with constant density on the set I_x with

$$\mu_{\beta,n}(I_x) = e^{-\beta V(x)},$$

where by I_x we denote the dyadic interval coded by $x = (x_1, \dots, x_n) \in \mathbb{T}_2$ such that $|x| = n$, i.e. $I_x = [\sum_{k=1}^n x_k 2^{-k}, \sum_{k=1}^n x_k 2^{-k} + 2^{-n})$.

After a normalization by proper a deterministic sequence, say $c_{\beta,n}$, one obtains measures $c_{\beta,n}\mu_{\beta,n}$ converging towards a finite nonzero random measure μ_β on I . Essentially, due to self-similarity of the model, asymptotic behavior of $\mu_{\beta,n}$ boils down to the asymptotic behavior of its total mass, i.e.

$$Z_{\beta,n} = \mu_{\beta,n}(I) = \sum_{|x|=n} e^{-\beta V(x)}.$$

Derrida and Spohn [20] explained that behavior of the cascade depends mainly on the parameter β and that there is a phase transition in the behavior of the limiting measure. Under (1.1) the critical value of parameter β is 1. For $\beta < 1$ (high temperature) and $\beta = 1$ (freezing temperature) the limiting measure μ_β is continuous, although singular with respect to the Lebesgue measure, whereas for $\beta > 1$ (low temperature) is purely atomic. In the continuous case one of the fundamental problems is description of local behavior of the measure μ_β , e.g. fluctuations of μ_β , which is the main problem considered in this paper. More precisely we aim to find optimal, deterministic functions ϕ_1 and ϕ_2 such that for μ_β -almost all $x \in I$ and for sufficiently large n we have almost surely (a.s.)

$$\phi_1(n) \leq \mu_\beta(B_n(x)) \leq \phi_2(n),$$

where $B_n(x)$ is the dyadic set of length 2^{-n} containing x .

1.2. The subcritical case. If $\beta < 1$ we say that the system is in the subcritical case or high temperature case. In this setting, result of Kahane and Peyrière [27] ensures that under mild integrability assumptions

$$(\mathbb{E}Z_{\beta,n})^{-1}Z_{\beta,n} \rightarrow Z_\beta \quad \text{a.s.}$$

where Z_β is a.s. positive and finite. Therefore one may infer, that for any fixed $x \in \mathbb{T}_2$, as $n \rightarrow \infty$

$$(\mathbb{E}Z_{\beta,n})^{-1}\mu_{\beta,n}(I_x) \rightarrow \mu_\beta(I_x) \quad \text{a.s.}$$

the details are given in Section 2. Local fluctuations of μ_β were described by Liu [29], who proved that for any $\varepsilon > 0$ for μ_β -almost any $x \in I$ and for sufficiently large n

$$e^{-(a+\varepsilon)n} \leq \mu_\beta(B_n(x)) \leq e^{-(a-\varepsilon)n},$$

for some constants $a > 0$ depending on β and ξ . Liu [29] proved this estimates for generalized model, defined in section 2.

1.3. The critical case. Whenever $\beta = 1$, we say that the system is in the critical case, or in the freezing temperature. This is the main case considered in this article. The situation is more involved than in the subcritical case since in the latter, the limit Z_β emerged as that of the positive martingale $(\mathbb{E}Z_{\beta,n})^{-1}Z_{\beta,n}$. Thus the proper choice of normalizing $c_{\beta,n}$ was natural. It runs out that in the critical case, this limit vanishes, showing that a different scaling is needed in order to obtain a nontrivial limit. The solution to this problem was recently delivered by Aïdékon and Shi [2] yielding

$$\sqrt{n}Z_{1,n} \rightarrow Z_1 \quad \text{in probability}$$

with a.s. finite and positive Z_1 . This convergence cannot be improved to a.s. convergence, since $\limsup_{n \rightarrow \infty} \sqrt{n}Z_{1,n} = \infty$ a.s. as is also proved in [2]. From the convergence in probability however, we obtain that, as $n \rightarrow \infty$

$$\sqrt{n}\mu_{1,n}(I_x) \rightarrow \mu_1(I_x) \quad \text{in probability.}$$

Barral et al. [4] proved that μ_1 is atomless and considered the problem of fluctuations. Under an additional assumption that ξ is Gaussian, it was proved that for certain $c > 0$ and arbitrary $k > 0$, with probability one for μ_1 -almost any $x \in I$, for sufficiently large n

$$e^{-c\sqrt{n \log n}} \leq \mu_1(B_n(x)) \leq e^{-k \log n}.$$

As the main results of this article shows, these bounds are not optimal and can be improved to the bounds of the form

$$e^{-d\sqrt{n \log \log n}} \leq \mu_1(B_n(x)) \leq e^{-\sqrt{n}L(n)},$$

for some slowly varying function L with $L(n) \rightarrow 0$ as $n \rightarrow \infty$. Moreover these estimates are valid for general, not necessary Gaussian, random variable ξ . Our results show also that these bounds are precise and give a detailed description of the lower and upper time-space envelope of $\log \mu_1(B_n(x))$. Roughly speaking, we will show that it satisfies the same bounds as the sequence $e^{-V(x_n)}$, where x_n is a vertex of n th generation chosen at random in a way that will allow us to describe the behavior of $V(x_n)$. For details see the discussion in Section 2.

1.4. The supercritical case. Situation when $\beta > 1$ is referred to as the supercritical case or "glassy" low temperature phase. In this case, the asymptotic behavior of $Z_{\beta,n}$ is determined by the minima of $V(x)$ for $|x| = n$. Using the work of Aïdékon [1], giving the weak convergence of $\min_{|x|=n} V(x) - 3/2 \log n$, Madaule [30] was able to prove that, as $n \rightarrow \infty$

$$n^{\frac{3\beta}{2}} Z_{\beta,n} \rightarrow Z_\beta \quad \text{in distribution.}$$

Whence we may infer that for $\beta > 1$

$$n^{\frac{3\beta}{2}} \mu_{\beta,n}(I_x) \rightarrow \mu_\beta(I_x) \quad \text{in distribution.}$$

However, as mentioned before, the limiting measure μ_β is purely atomic, see Barral et al. [8].

1.5. Liouville quantum gravity and Liouville Brownian motion. Finally, let us also mention that there is a natural counterpart of the measures μ_β in a continuous setting. Roughly speaking, these are random measures on a given domain D in \mathbb{R}^d with the following formal definition

$$\mu_\gamma = e^{\gamma X(x) - \gamma^2/2\mathbb{E}[X(x)^2]} \sigma(dx),$$

where X is a centered Gaussian random field with the property that $\text{Cov}(X(x), X(y)) = -\log \|x - y\| + O(1)$, as $\|x - y\| \rightarrow 0$ and σ is a given Radon measure. Since X is can not be defined pointwise even the existence of such a measure is far from being trivial (cf. [28, 10]). In the particular case, when X is a Gaussian free field with appropriate boundary condition these measures are important objects in theoretical physic because of the connections with the theory of Liouville quantum gravity. It is also worth to point out that Liouville quantum gravity is conjecturally related to discrete and continuum random surfaces. Roughly speaking, it appears as the scaling limit of a large class of statistical physical models (see for instance [34] for further details and references therein).

Recently, Berestycki [9] and independently by Garban, Rhodes and Vargas [25, 34] have constructed a diffusion on D , called Liouville Brownian motion, that has μ as a stationary measure. This process is conjectured to be the scaling limit of random walk on large planar maps. Finally, let us also mention the relationship between Liouville Brownian motion and the decay of the measure μ : if $\mathbf{p}_t(x, y)$ is its transition probability then it is believed that

$$\mathbf{p}_t(x, y) \sim \frac{e^{-\mu(B(x, |x-y|))/t}}{t} \quad \text{as } |x - y| \rightarrow 0.$$

See [34] for further discussion and details.

2. GENERALIZED MANDELBROT CASCADES AND MAIN RESULTS

2.1. Generalized Mandelbrot cascades. The main aim of this article is to study asymptotic properties of the limiting measure. Since only the values of the measure are of interest to us, we can regard the cascades as measures on some abstract space. This leads to so-called generalized Mandelbrot cascades defined in the next few paragraphs.

Consider a one-dimensional, discrete-time branching random walk governed by a point process Θ . We start with single particle placed at the origin of the real line. At time $n = 1$ this particle dies while giving birth to a random number of particles which will now form the first generation. Their position with respect to the birth place is determined by the point process Θ . At time $n = 2$ each particle of the first generation, independently from the rest, dies while giving birth to the particles from the second generation. Their positions, with respect to the birth place are determined by the same distribution Θ . The system goes on according to this rules. Obviously the number of particles in each generation forms a Galton-Watson process. We denote the corresponding random tree rooted at \emptyset by $\mathbb{T} \subseteq \mathbb{U}$, where

$$\mathbb{U} = \bigcup_{k \geq 0} \mathbb{N}^k.$$

We write $|x| = n$ if $x \in \mathbb{N}^n$, that is if x is a particle at n th generation. We denote the positions of particles of the n th generation as $(V(x) \mid |x| = n)$, and the whole process

as $(V(x) \mid x \in \mathbb{T})$. This process is usually referred to as a branching random walk. For any vertices $x, y \in \mathbb{T}$, by $\llbracket x, y \rrbracket$ we denote the shortest path connecting x and y . We can partially order \mathbb{T} by letting $x \geq y$ if $y \in \llbracket \emptyset, x \rrbracket$, that is if y is an ancestor of x . Let $x \wedge y = \inf\{x, y\}$ be the oldest common ancestor of x and y . Finally, x_i denotes the vertex in $\llbracket \emptyset, x \rrbracket$ such that $|x_i| = i$. The branching random walk gives rise to a random measure on the boundary of \mathbb{T} , i.e.

$$\partial\mathbb{T} = \{\xi \in \mathbb{I} \mid \xi_n \in \mathbb{T}, n \in \mathbb{N}\},$$

where $\mathbb{I} = \mathbb{N}^{\mathbb{N}}$. Notice that $\partial\mathbb{T}$ forms an ultrametric space with

$$B(x) = \{\xi \in \partial\mathbb{T} \mid \xi > x\}, \quad x \in \mathbb{T}$$

as its topological basis. This corresponds to the choice of $d_c(x, y) = c^{-|x \wedge y|}$ with $c > 1$ as a metric on $\partial\mathbb{T}$ in which $B(x)$ is a ball of radius $c^{-|x|}$.

Observe that if every particle has exactly two children, then $\mathbb{T} = \mathbb{T}_2$ and we are reduced to the Mandelbrot cascade defined in the Introduction, since the intervals $I_x \subseteq I$ correspond to the balls $B(x) \subseteq \partial\mathbb{T}$. However below we work in full generality, when the number of children, the corresponding tree \mathbb{T} and its boundary $\partial\mathbb{T}$ are random.

2.2. Assumptions and basic properties. In this paper we work under a standard assumption that the branching random walk is in the so-called boundary (or critical) case, that is

$$(2.1) \quad \mathbb{E} \left[\sum_{|x|=1} e^{-V(x)} \right] = 1 \quad \text{and} \quad \mathbb{E} \left[\sum_{|x|=1} V(x) e^{-V(x)} \right] = 0$$

which boils down to (1.1) if each particle has exactly two children. Throughout the paper we use the convention that $\sum_{\emptyset} = 0$. We need also some additional integrability assumptions, that is

$$(2.2) \quad \sigma^2 = \mathbb{E} \left[\sum_{|x|=1} V(x)^2 e^{-V(x)} \right] < \infty$$

as well as for some $p > 2$

$$(2.3) \quad \mathbb{E} [L(\log^+ L)^p] < \infty,$$

where $L = \sum_{|x|=1} (1 + V^+(x)) e^{-V(x)}$. Most of the discussions in this paper become trivial if the system dies out, that is if \mathbb{T} is finite. For example, by our definition $\partial\mathbb{T} = \emptyset$ for finite \mathbb{T} . Whence we assume that the underlying Galton-Watson process is supercritical, i. e. $\mathbb{E}[\sum_{|x|=1} 1] > 1$, so that the system survives with positive probability. Notice that this also implies the branching random walk is not reduced to a classical random walk, more precisely that the number of offspring $\#\Theta$ is bigger than 1 with positive probability. To avoid the need of considering the degenerate case we introduce the conditional probability

$$\mathbb{P}^*[\cdot] = \mathbb{P}[\cdot \mid \text{nonextinction}].$$

Our main results will be formulated in terms of the measure \mathbb{P}^* . We will now focus on the definition of the measures μ_n and μ starting with defining the total mass of the former

via

$$\mu_n(\partial\mathbb{T}) = \sum_{|x|=n} e^{-V(x)}.$$

It can be easily shown, that thanks to the first condition in (2.1) this sequence forms a non-negative, mean one martingale with respect to $\mathcal{F}_n = \sigma(V(x) \mid |x| \leq n)$ (called the *additive martingale*), and whence is convergent a.s. It turns out that our second assumption in (2.1) implies that the corresponding limit is 0 (see for example Biggins [12]). Nevertheless Aïdékon and Shi [2] proved that, under hypotheses (2.1), (2.2) and (2.3), we have the convergence

$$(2.4) \quad \sqrt{n}\mu_n(\partial\mathbb{T}) \rightarrow \mu(\partial\mathbb{T}) \quad \text{in probability.}$$

Moreover, $\mathbb{P}^*[\mu(\partial\mathbb{T}) > 0] = 1$. This result holds true and was proven under slightly weaker assumptions than (2.3). Since our main result requires (2.3), we will continue to invoke other results with slightly stronger conditions for readers convenience.

Similarly, to define $\mu_n(B(y))$ for $y \in \mathbb{T}$, just truncate the additive martingale to the subtree of all branches containing y , that is

$$\mu_n(B(y)) = \sum_{|x|=n, x>y} e^{-V(x)} = e^{-V(y)} \sum_{|x|=n, x>y} e^{-(V(x)-V(y))}$$

and by another appeal to (2.4) we infer that, as $n \rightarrow \infty$

$$\sqrt{n} \sum_{|x|=n, x>y} e^{-(V(x)-V(y))} \rightarrow W_y \quad \text{in probability}$$

for some nonnegative W_y . Whence, by defining $\mu(B(y))$ as the limit in probability of $\sqrt{n}\mu_n(B(y))$, we get

$$\mu(B(y)) = e^{-V(y)} W_y.$$

It can be easily verified that almost surely

$$\mu(\partial\mathbb{T}) = \sum_{|x|=1} e^{-V(x)} W_x.$$

Analogously, for any $y \in \mathbb{T}$ and $n > |y|$

$$W_y = \sum_{|x|=n, x>y} e^{-(V(x)-V(y))} W_x.$$

Note that this means exactly that

$$\mu\left(\bigcup_{|x|=n, x>y} B(x)\right) = \mu(B(y)) = \sum_{|x|=n, x>y} \mu(B(x))$$

since $B(y) = \bigcup_{|x|=n, x>y} B(x)$. By the extension theorem μ can be uniquely extended to a measure on $\partial\mathbb{T}$.

2.3. Main results. The main aim of this paper is to study local fluctuations of μ . For functions $f, g: \mathbb{N} \rightarrow \mathbb{R}$ we write $f(n) \leq_{i.o.} g(n)$ if $f(n) \leq g(n)$ for infinitely many n and $f(n) \leq_{a.a.} g(n)$ if $f(n) \leq g(n)$ for all but finitely many n . We want to find deterministic functions $\phi_1, \phi_2: \mathbb{N} \rightarrow \mathbb{R}$ such that

$$\phi_1(n) \leq_{a.a.} \mu(B(\xi_n)) \leq_{a.a.} \phi_2(n)$$

for μ -almost all $\xi \in \partial\mathbb{T}$, \mathbb{P}^* -almost surely.

Our first result describes the upper time-space envelope of $\mu(B(\xi_n))$.

Theorem 2.1. *Assume (2.1), (2.2) and (2.3). Let $\psi \in C^1(\mathbb{R}_+)$ be decreasing such that $t^{1/2-\delta}\psi(t)$ is increasing for some $\delta > 0$ and sufficiently large t . Then for μ -almost all $\xi \in \partial\mathbb{T}$, \mathbb{P}^* -almost surely*

$$\mu(B(\xi_n)) \leq_{a.a.} e^{-\sqrt{n}\psi(n)}, \quad \text{if } \int^\infty \frac{\psi(t)}{t} dt < \infty$$

and

$$\mu(B(\xi_n)) \geq_{i.o.} e^{-\sqrt{n}\psi(n)}, \quad \text{if } \int^\infty \frac{\psi(t)}{t} dt = \infty.$$

Note that this result gives necessary and sufficient conditions for $\mu(B(\xi_n)) \leq_{a.a.} e^{-\sqrt{n}\psi(n)}$ allowing to describe the upper time-space envelope of $\mu(B(\xi_n))$ with arbitrarily small gap. In order to illustrate this, define the functions $\psi_k, \psi_k^{(\varepsilon)}: \mathbb{N} \rightarrow \mathbb{R}$ for $k \in \mathbb{N}$ and $\varepsilon > 0$ by

$$\psi_k(t) = \left[\prod_{i=1}^k \log_{(i)}(t) \right]^{-1}, \quad \psi_k^{(\varepsilon)}(t) = \psi_k(t) \log_k(t)^{-\varepsilon},$$

where $\log_{(i)}(t)$ stands for the i th iterate of $\log(t)$. We can deduce from Theorem 2.1 that

$$e^{-\sqrt{n}\psi_k(n)} \leq_{i.o.} \mu(B(\xi_n)) \leq_{a.a.} e^{-\sqrt{n}\psi_k^{(\varepsilon)}(n)}.$$

Since the same inequalities would hold if $\psi_k^{(\varepsilon)}$ or ψ_k was multiplied by an arbitrary positive constant, we deduce that

$$\liminf_{n \rightarrow \infty} \frac{-\log(\mu(B(\xi_n)))}{\sqrt{n}\psi_k^{(\varepsilon)}(n)} = \infty \quad \text{and} \quad \liminf_{n \rightarrow \infty} \frac{-\log(\mu(B(\xi_n)))}{\sqrt{n}\psi_k(n)} = 0$$

for μ -almost all ξ , \mathbb{P}^* -almost surely. In particular for any $k \in \mathbb{N}$ one has

$$\liminf_{n \rightarrow \infty} \frac{\log(-\log(\mu(B(\xi_n)))) - \frac{1}{2} \log n + \sum_{j=2}^k \log_{(j)}(n)}{\log_{(k+1)}(n)} = -1.$$

Our second result describes the lower time-space envelope.

Theorem 2.2. *Assume (2.1), (2.2), (2.3) and $\mathbb{E} \left[\sum_{|x|=1} |V(x)|^{3+\varepsilon} e^{-V(x)} \right] < \infty$ for some $\varepsilon > 0$. Then for any $\delta > 0$ and μ -almost all $\xi \in \partial\mathbb{T}$, \mathbb{P}^* -a.s.*

$$\mu(B(\xi_n)) \geq_{a.a.} e^{-(1+\delta)\sqrt{2\sigma^2 n \log \log n}},$$

where σ^2 is given by (2.2) and

$$\mu(B(\xi_n)) \leq_{i.o.} e^{-(1-\delta)\sqrt{2\sigma^2 n \log \log n}}.$$

From the above one gets instantly that

$$\limsup_{n \rightarrow \infty} \frac{-\log(\mu(B(\xi_n)))}{\sqrt{2\sigma^2 n \log \log n}} = 1 \text{ a.s.}$$

2.4. Discussion of the results. The problem of describing local fluctuations of the Mandelbrot cascades in the critical case was previously investigated by Barral et al. [4]. They considered the case when \mathbb{T} is the binary tree \mathbb{T}_2 and the branching random walk is generated by the Gaussian distribution and proved that for any $\varepsilon > 0$ and any $k \in \mathbb{N}$

$$\exp \left\{ -(1 + \varepsilon) \sqrt{2 \log(2) n \log n} \right\} \leq_{i.o.} \mu(B(\xi_n)) \leq_{a.a.} \exp \{-k \log n\}$$

and

$$\exp \left\{ -\sqrt{6 \log(2)} \sqrt{n(\log n + (1/3 + \varepsilon) \log \log n)} \right\} \leq_{a.a.} \mu(B(\xi_n))$$

for μ -almost all $\xi \in \partial \mathbb{T}$, \mathbb{P}^* -almost surely. Notice that the bounds appearing in the first part of this result, have different asymptotic. Thus, this result does not give a detailed information about the upper time-space envelope of $\mu(B(\xi_n))$.

To prove Theorems 2.1 and 2.2 we use the spinal decomposition and the change of measure, based on the work of Biggins, Kyprianou [13], used for example by Aïdékon, Shi [2]. After the change of measure, along every spine $\{w_n\}$ (i. e. a random element of $\partial \mathbb{T}$), in the new probability space, $V(w_n)$ behaves as a random walk conditioned to stay positive and its fluctuations were studied by Hambly et al. [26]. This result provides description of time-space envelopes with arbitrary small error.

The details are given in Sections 3 – 6. First, in Section 3, we recall some basic properties of the random walk conditioned to stay positive. In Section 4 we describe the change of the probability space and finally in Sections 5 and 6 we give complete proof of our results.

Let us finally mention that similar results can be obtained for the critical Gaussian multiplicative chaos. Then, the problem of local fluctuations was considered by Barral et al. [5]. Based on techniques described in Duplantier et al. [21, 22] one can prove analogous result to Theorems 2.1 and 2.2. This is the subject of a work in progress.

3. ONE-DIMENSIONAL RANDOM WALK

In this section we introduce a one-dimensional random walk associated with the branching random walk defined above. Next we define a random walk conditioned to stay above some level $-\alpha$ for $\alpha \geq 0$ and formulate its fundamental properties concerning fluctuations of its paths. Those results will play a crucial role in our arguments.

3.1. An associated one-dimensional random walk. Assumption (2.1) allows us to introduce a random walk $\{S_n\}$ with the distribution of increments given by

$$\mathbb{E}[f(S_1)] = \mathbb{E} \left[\sum_{|x|=1} f(V(x)) e^{-V(x)} \right],$$

for any measurable $f: \mathbb{R} \rightarrow [0, \infty)$. Then, since the increments $S_2 - S_1, S_3 - S_2, \dots$ are independent copies of S_1 , one can easily show, that for any $n \in \mathbb{N}$ and measurable $g: \mathbb{R}^n \rightarrow \mathbb{R}$

we have

$$\mathbb{E}[g(S_1, S_2, \dots, S_n)] = \mathbb{E} \left[\sum_{|x|=n} g(V(x_1), V(x_2), \dots, V(x_n)) e^{-V(x)} \right].$$

Note that by (2.1)

$$\mathbb{E}[S_1] = \mathbb{E} \left[\sum_{|x|=1} V(x) e^{-V(x)} \right] = 0$$

and thus the random walk $\{S_n\}$ is centered and by (2.2) has a finite variance

$$\mathbb{E}[S_1^2] = \mathbb{E} \left[\sum_{|x|=1} V(x)^2 e^{-V(x)} \right] = \sigma^2 < \infty.$$

3.2. A conditioned random walk. It turns out that in our considerations an important role will be played not by the random walk $\{S_n\}$, but by its trajectories conditioned to stay above $-\alpha$ for some $\alpha \geq 0$. Bertoin and Doney [11] showed that for each $A \in \sigma(S_j, j \leq k)$ the limiting probabilities

$$\lim_{n \rightarrow \infty} \mathbb{P}[A | \tau_\alpha > n],$$

where $\tau_\alpha = \inf\{k \geq 1 : S_k < -\alpha\}$, are well defined and nontrivial. Their result is a discrete analogue of the relationship between the Brownian motion and the Bessel-3 process. It turns out that the conditioned random walk forms a Markov chain. Here we sketch the arguments leading to a description of its transition probability.

Since the random walk $\{S_n\}$ is centered, $\tau^+ = \inf\{k \geq 1 : S_k \geq 0\}$ is finite a.s. If we put

$$(3.1) \quad R(u) = \mathbb{E} \left[\sum_{j=0}^{\tau^+-1} \mathbb{1}_{\{S_j \geq -u\}} \right],$$

we see that by the duality lemma, R is the renewal function associated with the entrance of $(-\infty, 0)$ by the walk S . That is, R can be written in the following fashion

$$R(u) = \sum_{k \geq 0} \mathbb{P}[H_k \geq -u]$$

for $u \geq 0$, where $H_0 > H_1 > H_2 > \dots$ are strictly descending leader heights of $\{S_n\}$, $H_k = S_{\tau_k^-}$ with $\tau_0^- = 0$ and $\tau_{k+1}^- = \min\{j \geq \tau_k^- : S_j < H_k\}$ for $k \geq 0$. One can show that $\mathbb{E}[S_1^2] < \infty$ and $\mathbb{E}[S_1] = 0$ ensure $\mathbb{E}[|H_1|] < \infty$ (see e.g. Feller [24], Theorem XVIII.5.1).

As a consequence of (3.1), by conditioning on S_1 , one gets the following identity

$$(3.2) \quad R(u) = \mathbb{E}[R(S_1 + u) \mathbb{1}_{\{S_1 \geq -u\}}] \quad \text{for } u \geq 0.$$

Thus $\mathbb{1}_{\{\tau_\alpha > n\}} R(S_n + \alpha)$ is a martingale. The corresponding h -transform defines a Markov chain such that for any measurable subset A of \mathbb{R}^k

$$(3.3) \quad \mathbb{P}_\alpha^\uparrow[(S_1, \dots, S_k) \in A] = \frac{1}{R(\alpha)} \mathbb{E}[\mathbb{1}_{\{(S_1, \dots, S_k) \in A\} \cap \{\tau_\alpha > k\}} R(S_k + \alpha)].$$

Then the transition probability of this Markov chain is given by

$$P_\alpha^\uparrow(x, dy) = \mathbb{1}_{\{y \geq \alpha\}} \frac{R(y + \alpha)}{R(x + \alpha)} P(S_1 + x \in dy), \quad x \geq -\alpha.$$

The random process $\{S_n\}$ under the probability measure $\mathbb{P}_\alpha^\uparrow$ is called the random walk conditioned to stay in $[-\alpha, \infty)$.

3.3. Some properties of the renewal function R . Here we collect some properties of the function R , following from the renewal theorem, that will be needed in next sections. The renewal theorem (see e.g. Feller [24]) distinguishes between two cases, when the random walk $\{S_n\}$ is nonarithmetic (i.e. it is not contained in any set of the form $a\mathbb{Z}$ for positive a) and when it is arithmetic. In the first case the renewal theorem says that for every $h > 0$ the limit

$$\lim_{u \rightarrow \infty} R(u+h) - R(u) = h/\mathbb{E}|H_1|$$

exists and is finite. If the random walk is arithmetic, then the same limit exists but only for h and u being multiples of a :

$$\lim_{n \rightarrow \infty} R(na+h) - R(na) = h/\mathbb{E}|H_1|.$$

Below we treat both cases simultaneously since we need just some simple consequences of the results stated above. In both cases the following limit exists

$$(3.4) \quad c_0 = \lim_{u \rightarrow \infty} \frac{R(u)}{u}.$$

Whence there are constants $c_2 > c_1 > 0$ such that for any $u \geq 0$

$$(3.5) \quad c_1(1+u) \leq R(u) \leq c_2(1+u).$$

Moreover, there is a constant $c_3 > 0$ such that for every $u, x > 0$

$$(3.6) \quad R(u+x) - R(u) \leq c_3(1+x).$$

3.4. Some properties of the conditioned random walk. Here we describe some properties of trajectories of the Markov chain $(\{S_n\}, \mathbb{P}_\alpha^\uparrow)$ that will be needed in the proofs of our main results. Analogously to the Bessel-3 process, paths of the conditioned random walk stays in 'some neighborhood' of $n^{1/2}$. The precise description of its fluctuations was provided by Hambly et al. [26] and is stated in the next two lemmas.

Lemma 3.1 (Law of the iterated logarithm). *Suppose that for some $\delta > 0$, $\mathbb{E}|S_1|^{3+\delta} < \infty$. Then*

$$\limsup_{n \rightarrow \infty} \frac{S_n}{\sqrt{2n\sigma^2 \log \log n}} = 1 \quad \mathbb{P}_\alpha^\uparrow \text{ a.s.}$$

Lemma 3.2 (Lower Space-Time Envelope). *Suppose that $\mathbb{E}[S_1^2] < \infty$ and ψ is a function on $(0, \infty)$ such that $\psi(t) \downarrow 0$ and $\sqrt{t}\psi(t) \uparrow \infty$ as $t \uparrow \infty$. Then*

$$\liminf_{n \rightarrow \infty} \frac{S_n}{\sqrt{n}\psi(n)} = \infty \text{ or } 0 \quad \mathbb{P}_\alpha^\uparrow \text{ a.s.}$$

accordingly as

$$\int_0^\infty \frac{\psi(t)}{t} dt < \infty \text{ or } = \infty.$$

We will need also two further auxiliary lemmas reflecting the fact that trajectories of the conditioned random walk goes to $+\infty$. The first lemma is due to Biggins [14].

Lemma 3.3. *Fix $y \geq x \geq -\alpha$. Then*

$$\mathbb{P}_\alpha^\uparrow \left[\min_{n \geq 1} S_n > x \mid S_0 = y \right] = \frac{R(y-x)}{R(\alpha+y)}.$$

The next lemma seems to be standard, however since we don't know any reference we provide a complete proof of it.

Lemma 3.4. *For fixed $x > 0$ there is c_4 such that*

$$\mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x \right] \leq \frac{c_4 \log n}{\sqrt{n}}, \quad n > 1.$$

Proof. In the proof we need the local limit theorem for conditioned random walks due to Caravenna [19]. In our settings this result implies that for fixed $h > 0$ there is c_5 such that

$$(3.7) \quad \sup_{r \geq -\alpha} \mathbb{P}_\alpha^\uparrow [r \leq S_n \leq r+h] \leq \frac{c_5}{\sqrt{n}},$$

for any $n \geq 1$. We write

$$(3.8) \quad \begin{aligned} & \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x \right] \\ & \leq \mathbb{P}_\alpha^\uparrow [S_n \leq 2x] + \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x; S_n > nx \right] + \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x; 2x < S_n \leq nx \right]. \end{aligned}$$

Thus we have to bound three expressions. The first term, by (3.7), can be bounded in the following fashion:

$$\mathbb{P}_\alpha^\uparrow [S_n \leq 2x] \leq \frac{c_5(2x+\alpha)}{h\sqrt{n}}.$$

For the second one we use Lemma 3.3, inequality (3.6) and the lower bound in (3.5)

$$\begin{aligned} \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x; S_n > nx \right] &= \int_{nx}^\infty \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x \mid S_n = y \right] \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\ &= \int_{nx}^\infty \left(1 - \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k > x \mid S_n = y \right] \right) \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\ &= \int_{nx}^\infty \left(1 - \frac{R(y-x)}{R(\alpha+y)} \right) \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\ &= \int_{nx}^\infty \frac{R(\alpha+y) - R(y-x)}{R(\alpha+y)} \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\ &\leq \frac{c_3}{c_1} \int_{nx}^\infty \frac{1+\alpha+x}{1+\alpha+y} \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\ &\leq \frac{c_3}{c_1} \frac{1+\alpha+x}{1+\alpha+nx} \mathbb{P}_\alpha^\uparrow [S_n > nx] \\ &\leq \frac{c_6}{n}, \end{aligned}$$

for some c_6 . To estimate the last term in (3.8) we apply, successively, Lemma 3.3, (3.6), the lower bound in (3.5) and (3.7)

$$\begin{aligned}
& \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x; 2x < S_n \leq nx \right] \\
&= \int_{2x}^{nx} \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k \leq x | S_n = y \right] \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\
&= \int_{2x}^{nx} \left(1 - \mathbb{P}_\alpha^\uparrow \left[\min_{k \geq n} S_k > x | S_n = y \right] \right) \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\
&= \int_{2x}^{nx} \frac{R(\alpha + y) - R(x - y)}{R(\alpha + y)} \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\
&\leq \frac{c_3}{c_1} \int_{2x}^{nx} \frac{1 + \alpha + x}{1 + \alpha + y} \mathbb{P}_\alpha^\uparrow [S_n \in dy] \\
&\leq \frac{c_3}{c_1} \sum_{0 \leq i \leq xn/h} \frac{1 + \alpha + x}{1 + \alpha + 2x + ih} \mathbb{P}_\alpha^\uparrow [2x + ih \leq S_n \leq 2x + (i + 1)h] \\
&\leq \frac{c_7(1 + \log(n))}{\sqrt{n}}.
\end{aligned}$$

This completes the proof. \square

4. DERIVATIVE MARTINGALE AND CHANGE OF PROBABILITIES

4.1. Derivative martingale. To study local properties of the random measure μ it is convenient to express it in terms of another fundamental martingale associated with the branching random walk, namely of the derivative martingale. It is defined as

$$D_n = \sum_{|x|=n} V(x) e^{-V(x)}.$$

Our assumption (2.1) ensures that this formula defines a centered martingale, i.e. $\mathbb{E} D_n = 0$ for all $n \in \mathbb{N}$. Convergence of the derivative martingale was studied by Biggins and Kyprianou [13], who proved that under assumptions (2.1), (2.2) and (2.3)

$$D_n \rightarrow D, \quad \mathbb{P}^* \text{ a.s.}$$

and $D > 0$, \mathbb{P}^* a.s. Aïdékon and Shi [2] were able to relate D with the limit of the additive martingale, that is $\mu(\partial\mathbb{T})$ (see (2.4)) and proved that

$$\mu(\partial\mathbb{T}) = \left(\frac{2}{\pi\sigma^2} \right)^{1/2} D \quad \mathbb{P}^* \text{ a.s.}$$

Similarly, starting the derivative martingale from any vertex $x \in \mathbb{T}$, that is considering

$$D_{x,n} = \sum_{\substack{|y|=|x|+n \\ y>x}} (V(y) - V(x)) e^{-(V(y)-V(x))},$$

gives a.s. limit $D_x = \lim_{n \rightarrow \infty} D_{x,n}$. Since $W_x = c_8 D_x$ we get

$$(4.1) \quad \mu(B(x)) = c_8 e^{-V(x)} D_x \quad \mathbb{P}^* \text{ a.s.},$$

where $c_8 = \left(\frac{2}{\pi\sigma^2}\right)^{1/2}$.

4.2. Change of probabilities. Our argument for Theorems 2.1 and 2.2 require a change of probability space. This is a standard approach in the theory of branching random walks. An appropriate change of the probability measure reduces the main problem to a question expressed in terms of a random walk on the real line. We would like to use the fact that $\{D_n\}$ is a martingale and apply the Doob h-transform. Unfortunately the derivative martingale is not positive. To overcome this difficulty we follow the approach based on the truncated argument presented in Biggins, Kyprianou [13] and Adkon, Shi [2]. For any vertex $x \in \mathbb{T}$ put

$$\underline{V}(x) = \min_{y \in \llbracket \emptyset, x \rrbracket} V(y).$$

Define the truncated martingale as

$$(4.2) \quad D_n^{(\alpha)} = \sum_{|x|=n} R(V(x) + \alpha) e^{-V(x)} \mathbb{1}_{\{\underline{V}(x) \geq -\alpha\}},$$

where R is given by (3.1). Because of (2.1) and (3.4), we expect that for large values of α , $D_n^{(\alpha)}$ should be comparable with D_n . In next sections we describe how these martingales are related with each other in terms of the cascade measures.

Assuming (2.1), Biggins and Kyprianou [13] proved that for any $\alpha \geq 0$, $\{D_n^{(\alpha)}\}$ is a nonnegative martingale with $\mathbb{E}[D_n^{(\alpha)}] = R(\alpha)$. Using this fact we can define a probability measure $\mathbb{P}^{(\alpha)}$ via

$$\mathbb{P}^{(\alpha)}|_{\mathcal{F}_n} = \frac{D_n^{(\alpha)}}{R(\alpha)} \cdot \mathbb{P}|_{\mathcal{F}_n}$$

that is for any $A \in \mathcal{F}_n$,

$$(4.3) \quad \mathbb{P}^{(\alpha)}[A] = \mathbb{E}\left[\mathbb{1}_A \frac{D_n^{(\alpha)}}{R(\alpha)}\right].$$

Giving a rigorous arguments requires $\mathbb{P}^{(\alpha)}$ to be defined on the space of marked trees with distinguished rays, that is infinite lines of descents starting from the root or simply the elements of $\partial\mathbb{T}$ (we refer to Neveu [33] for more details). The distinguished ray is called the spine and will be denoted by $\{w_n\}$. To explain how it is chosen, let for any $u > \alpha$, $\hat{\Theta}^{(\alpha)}(u)$ denote a point process whose distribution under \mathbb{P} is the law of $(u + V(x), |x| = 1)$ under $\mathbb{P}^{(\alpha+u)}$. We start with a single particle placed at the origin of the real line. Denote this particle by $w_0 = \emptyset$. At n th moment in time (for $n > 0$), each particle of generation n dies and gives birth to point processes independently of each other: the particle w_n generates a point process distributed as $\hat{\Theta}^{(\alpha)}(V(w_n))$ whereas other particle say x , with $|x| = n$ and $x \neq w_n$ generates a point process distributed as $V(x) + \Theta$. Finally the particle w_{n+1} is chosen among the children y of w_n with probability proportional to $R(\alpha + V(y))e^{-V(y)} \mathbb{1}_{\{\underline{V}(y) \geq -\alpha\}}$. An induction argument proves

$$\mathbb{P}^{(\alpha)}[w_n = x | \mathcal{F}_n] = \frac{R(V(x) + \alpha) e^{-V(x)} \mathbb{1}_{\{\underline{V}(x) \geq -\alpha\}}}{D_n^{(\alpha)}}.$$

Note that, formally, the measure defined above is different than the one defined by the equation (4.3). However, since there is a natural projection from the space of marked trees with distinguished rays to the space of marked trees and (4.3) defines marginal law of $\mathbb{P}^{(\alpha)}$ defined on space of marked trees with distinguished rays, we feel free, by slight abuse of notation, to use the same symbol for both measures.

4.3. Spine and conditioned random walk. Biggins and Kyprianou [13] proved that the positions of the particles obtained in the way described above have the same distribution as the branching random walk under $\mathbb{P}^{(\alpha)}$. Moreover the process $\{V(w_n)\}$ under $\mathbb{P}^{(\alpha)}$, is distributed as the centered random walk $\{S_n\}$ conditioned to stay in $[-\alpha, \infty)$.

Since the truncated martingale (4.2) is positive, it has an a.s. limit

$$(4.4) \quad D^{(\alpha)} = \lim_{n \rightarrow \infty} D_n^{(\alpha)}.$$

It turns out (see Biggins and Kyprianou [13]) that this convergence holds also in mean. This implies in particular that $\mathbb{P}^{(\alpha)}$ is absolutely continuous with respect to \mathbb{P} with density $D^{(\alpha)}$, that is for any $A \in \mathcal{F}$

$$\mathbb{P}^{(\alpha)}[A] = \mathbb{E} \left[\mathbb{1}_A \frac{D^{(\alpha)}}{R(\alpha)} \right].$$

5. SOME PROPERTIES OF THE RANDOM MEASURE μ AND AUXILIARY LEMMAS

In this section we are going to prove some further properties of μ that will be needed in the proofs of Theorems 2.1 and 2.2. Our main aim is to obtain a formula for the measure μ more suitable for our needs. This will be done in Lemma 5.4. However, to achieve this point we need a number of purely technical steps. We reduce the probability measure to $\mathbb{P}^{(\alpha)}$. Then we introduce truncated cascades, which provides an alternative formula for μ (see (5.4) below) on a set of large measure. This approach allows us to control the measure of a ball near a typical point and, in particular, prove its continuity needed in Lemma 5.4. The fact that μ is continuous, under stronger assumptions, has been already shown in [4], however here we work under much weaker moment hypotheses and cannot apply directly that results.

5.1. Reduction to the measure $\mathbb{P}^{(\alpha)}$. Our main result concerns \mathbb{P}^* -a.s. fluctuations of $\mu(B(\xi_n))$ along infinite path $\xi \in \partial\mathbb{T}$ that does not belong to the null set of a measure μ . However, instead of sampling ξ according to μ (or its normalized version) we will take it as the random spine $\{w_n\}$. We already know that for $\alpha \geq 0$ the process $(\mathbb{P}^{(\alpha)}, \{V(w_n)\})$ behaves like a conditioned random walk. However, in order to reduce the main problem to this setting, we need to know that μ -a.e. element from $\partial\mathbb{T}$ belongs to the range of a spine in $(\mathbb{P}^{(\alpha)}, \{w_n\})$ for some parameter α . In other words we need to prove that the range of the spines $(\mathbb{P}^{(\alpha)}, \{w_n\})$ is a relatively big subset of $(\mathbb{P}, \partial\mathbb{T})$. This is explained in the following Lemma.

Lemma 5.1. *Assume (2.1) and (2.3). We have*

$$\mathbb{P}^*[D^{(\alpha)} = 0] \leq c_9 e^{-\alpha} \rightarrow 0 \quad \text{as } \alpha \rightarrow \infty.$$

Proof. By (3.5)

$$R(x + \alpha) \geq c_2(1 + x + \alpha) \geq c_2x.$$

Since $D > 0$, \mathbb{P}^* -a.s., on the set $\{\min_{x \in \mathbb{T}} V(x) \geq -\alpha\}$ we have $D^{(\alpha)} \geq c_2D > 0$. Therefore,

$$\mathbb{P}^*[D^{(\alpha)} = 0] \leq \mathbb{P}^*\left[\min_{x \in \mathbb{T}} V(x) < -\alpha\right] \leq c_9e^{-\alpha},$$

(see inequality (2.2) in [18]). \square

The above lemma reduces our problem to describing fluctuations of the sequence $\mu(B(w_n))$ under the measure $\mathbb{P}^{(\alpha)}$. By (4.1)

$$(5.1) \quad \mu(B(w_n)) = c_8e^{-V(w_n)}D_{w_n}.$$

The main idea is to show that the growth of D_{w_n} does not interfere in the behavior of $\mu(B(w_n))$ that should be governed by $e^{-V(w_n)}$. However, under the changed measure $\mathbb{P}^{(\alpha)}$ the law of D_{w_n} depends on $V(w_n)$: conditioned on $V(x)$ and $V(w_k)$ for $|x|, k \leq n$, D_{w_n} under $\mathbb{P}^{(\alpha)}$ has the same law as D under $\mathbb{P}^{(\alpha+V(w_n))}$ and it can be easily seen that the sequence is not even tight. This problem is the most significant difference between the critical and subcritical case and in order to overcome this issue we need a convenient representation of $\mu(B(w_n))$, which is available in a slightly different settings.

Notice that, by Lemma 5.1, if some property holds with probability $\mathbb{P}^{(\alpha)}$ equal to 1, then this property must hold with probability \mathbb{P}^* at least $1 - c_9e^{-\alpha}$. If α can be taken arbitrarily big, this implies that the property in question also hold with probability \mathbb{P}^* equal to 1. Thus, it is sufficient to prove estimates stated in Theorems 2.1 and 2.2 for $\mathbb{P}^{(\alpha)}$ a.e. spine $\{w_n\}$ and all $\alpha \geq 0$ sufficiently large.

5.2. Truncated cascades. The truncated martingale $D_n^{(\alpha)}$ introduced in the previous section is a useful tool to provide a different construction of the measure μ . The idea is to define a truncated version of Mandelbrot cascades that converge to some limit measure which with high probability, up to a constant, coincide with μ . The advantage of this approach is that it allows us to prove the upper bound in Theorem 2.1 and deduce continuity of measure μ .

For given $\alpha \geq 0$ and any $x \in \mathbb{T}$ we consider the martingale

$$(5.2) \quad D_{x,n}^{(\alpha)} = \sum_{\substack{|y|=|x|+n \\ y>x}} R(V(y) + \alpha)e^{-(V(y)-V(x))} \mathbb{1}_{\{\underline{V}^x(y) \geq -\alpha\}},$$

where for $y > x$, $\underline{V}^x(y) = \min_{z \in \llbracket x, y \rrbracket} V(z)$. As before $D_{x,n}^{(\alpha)}$ converges almost surely and in the mean to the limit $D_x^{(\alpha)}$. We may define now the measure $\mu^{(\alpha)}$ on $\partial\mathbb{T}$ by setting

$$(5.3) \quad \mu^{(\alpha)}(B(x)) = \mathbb{1}_{\{\underline{V}(x) \geq -\alpha\}} e^{-V(x)} D_x^{(\alpha)}.$$

Note that since $\mu_n(\partial\mathbb{T}) = \sum_{|x|=n} e^{-V(x)} \rightarrow 0$, $\inf_{|x|=n} V(x) \rightarrow \infty$. Thus, by (3.4) we have

$$\mu^{(\alpha)}(B(x)) = c_0e^{-V(x)}D_x$$

on the set $\{\min_{x \in \mathbb{T}} V(x) > -\alpha\}$. Since the probability of the last event is at least $1 - c_9 e^{-\alpha}$ (cf. formula (2.2) in [18]) we have

$$(5.4) \quad \mu = \frac{c_0}{c_8} \mu^{(\alpha)},$$

with probability at least $1 - c_9 e^{-\alpha}$.

5.3. Continuity of μ and upper estimates. The goal of this subsection is to establish that if $\int^\infty \psi(t) t^{-1} dt < \infty$, then

$$\mu(B(w_n)) \leq_{a.a.} e^{-\sqrt{n}\psi(n)} \quad \mathbb{P}^* \text{-a.s.}$$

This will in particular imply that μ is continuous. As explained above we deduce this result from analogous properties of the measure $\mu^{(\alpha)}$ considered with respect to $\mathbb{P}^{(\alpha)}$ for arbitrary large α . Therefore it follows immediately from the Lemma:

Lemma 5.2. *Under hypotheses of Theorem 2.1*

$$\mu^{(\alpha)}(B(w_n)) \leq_{a.a.} e^{-\sqrt{n}\psi(n)} \quad \mathbb{P}^{(\alpha)} \text{-a.s.}$$

Proof. Step 1. First we prove that for any n

$$(5.5) \quad \mu^{(\alpha)}(B(w_n)) \leq \sum_{k \geq n} e^{-V(w_k)} \sup_{l \geq 1} D_{w_k, l}^{(\alpha)}, \quad \mathbb{P}^{(\alpha)} \text{-a.s.}$$

Indeed, from the definition (5.2) of $D_{x, n}^{(\alpha)}$ it follows that for any $N > 0$

$$e^{-V(w_n)} D_{w_n, N}^{(\alpha)} = e^{-V(w_{n+1})} D_{w_{n+1}, N-1}^{(\alpha)} + \sum_{x \in C(w_n) \setminus \{w_{n+1}\}} e^{-V(x)} D_{x, N-1}^{(\alpha)},$$

where $C(w_n) = \{|x| = n+1, x > w_n\}$ denotes the set of children of w_n . Iterating this equation yields

$$\begin{aligned} e^{-V(w_n)} D_{w_n, N}^{(\alpha)} &= R(V(w_{n+N}) + \alpha) e^{-V(w_{n+N})} + \sum_{k=0}^{N-1} \sum_{x \in C(w_{n+k}) \setminus \{w_{n+k+1}\}} e^{-V(x)} D_{x, N-k-1}^{(\alpha)} \\ &\leq R(V(w_{n+N}) + \alpha) e^{-V(w_{n+N})} + \sum_{k=0}^{N-1} \sup_{l \geq 0} \sum_{x \in C(w_{n+k})} e^{-V(x)} D_{x, l}^{(\alpha)}. \end{aligned}$$

Since $V(w_{n+N}) \rightarrow \infty$, $\mathbb{P}^{(\alpha)}$ a.s., letting N tend to infinity we get

$$e^{-V(w_n)} D_{w_n}^{(\alpha)} \leq \sum_{k \geq n} \sup_{l \geq 0} \sum_{x \in C(w_k)} e^{-V(x)} D_{x, l}^{(\alpha)} = \sum_{k \geq n} e^{-V(w_k)} \sup_{l \geq 1} D_{w_k, l}^{(\alpha)}.$$

Thus, in view of (5.3) we have (5.5).

Step 2. Next we prove that the contribution of $\sup_{l \geq 1} D_{w_k, l}^{(\alpha)}$ in (5.5) is negligible. For this purpose we show that for $\delta > 1/p$ (for p defined in (2.3)) we have

$$(5.6) \quad G_n^{(\alpha)} := \frac{1}{R(\alpha)} \sup_{l \geq 1} D_{w_n, l}^{(\alpha)} \leq_{a.a.} e^{n^\delta} \quad \mathbb{P}^{(\alpha)} \text{-a.s.}$$

To prove this inequality first we show that for $\beta > 0$ and $F(\beta) := \mathbb{E}^{(\beta)}[(\log^+(G_0^{(\beta)}))^p]$ we have

$$(5.7) \quad \sup_{\beta > 0} F(\beta) = c_{10} < \infty.$$

By (3.5) we can take constant $c_{11} > 0$ such that for all $\beta > 0$ and $x \geq -\beta$ we have

$$\frac{R(\beta + x)}{R(\beta)} \leq c_{11}(1 + x^+).$$

Then, by the definition of the measure $\mathbb{P}^{(\beta)}$, we can write

$$\begin{aligned} \mathbb{E}^{(\beta)} \left[\left(\log^+ (G_0^{(\beta)}) \right)^p \right] &\leq \mathbb{E} \left[\left(\log^+ \left(\sup_n D_n^{(\beta)} / R(\beta) \right) \right)^p \sum_{|y|=1} \frac{R(V(y) + \beta)}{R(\beta)} e^{-V(y)} \mathbb{1}_{\{V(y) \geq -\beta\}} \right] \\ &\leq (2p)^p c_{11} \mathbb{E} \left[\left(\log^+ \left(\sup_n D_n^{(\beta)} / R(\beta) \right)^{1/2p} \right)^p \sum_{|y|=1} (1 + V^+(y)) e^{-V(y)} \right]. \end{aligned}$$

For the latter we can use the simple inequality

$$ab \leq e^a + b \log^+ b,$$

valid for any $a, b \geq 0$. Taking

$$a = \log^+ \left(\sup_n D_n^{(\beta)} / R(\beta) \right)^{1/2p} \quad \text{and} \quad b^p = L = \sum_{|y|=1} (1 + V^+(y)) e^{-V(y)}$$

the above expectation can be bounded in the following way

$$\begin{aligned} \mathbb{E}^{(\beta)} \left[\left(\log^+ (G_0^{(\beta)}) \right)^p \right] &\leq \mathbb{E} \left[\left(\left(\sup_n D_n^{(\beta)} / R(\beta) \right)^{1/2p} + 1 + \frac{1}{p} L^{1/p} \log^+ L \right)^p \right] \\ &\leq 3^p \mathbb{E} \left[\left(\sup_n D_n^{(\beta)} / R(\beta) \right)^{1/2} + 1 + L (\log^+ L)^p \right] < \infty, \end{aligned}$$

since by Doob's martingale inequality $\mathbb{E} \left[\left(\sup_n D_n^{(\beta)} / R(\beta) \right)^{1/2} \right] \leq 2$. This proves (5.7).

Then for any $\alpha \geq 0$ and $n \in \mathbb{N}$ we have

$$\begin{aligned} \mathbb{P}^{(\alpha)} \left[G_n^{(\alpha)} > e^{n^\delta} \right] &= \mathbb{E}^{(\alpha)} \left[\mathbb{P}^{(\alpha)} \left[G_n^{(\alpha)} > e^{n^\delta} \mid V(w_n) \right] \right] \\ &= \mathbb{E}^{(\alpha)} \left[\mathbb{P}^{(\alpha)} \left[\left(\log^+ (G_n^{(\alpha)}) \right)^p > n^{p\delta} \mid V(w_n) \right] \right] \\ &\leq \mathbb{E}^{(\alpha)} \left[F(\alpha + V(w_n)) n^{-p\delta} \right] \leq c_{10} n^{-p\delta}. \end{aligned}$$

The claim follows by the Borel-Cantelli lemma.

Step 3. Now we prove the required upper bound. Recall that $(\mathbb{P}^{(\alpha)}, \{w_n\})$ has the same distribution as the conditioned random walk $(\mathbb{P}_\alpha^\uparrow, S_n)$. We bound the $\mu^{(\alpha)}$ using Lemma 3.2 and (5.6). Let us take any ψ such that $\int^\infty \psi(t)t^{-1} dt < \infty$, then

$$V(w_n) >_{a.a.} 4\sqrt{n}\psi(2n) \quad \mathbb{P}^{(\alpha)} \text{ a.s.}$$

and hence, by (5.5) and (5.6)

$$\mu^{(\alpha)}(B(w_n)) \leq_{a.a.} \sum_{j \geq n} e^{-\sqrt{j}3\psi(2j)+j^\delta} \quad \mathbb{P}^{(\alpha)}\text{-a.s.}$$

Since the function $t^{1/2-\delta}\psi(t)$ is increasing, we have for sufficiently large n

$$\sum_{j \geq n} e^{-3\sqrt{j}\psi(2j)+j^\delta} \leq \sum_{j \geq n} e^{-2\sqrt{j}\psi(2j)}.$$

Thus, to prove the result it is sufficient to justify that

$$(5.8) \quad \sum_{j \geq n} e^{-2\sqrt{j}\psi(2j)} \leq e^{-\sqrt{n}\psi(n)}.$$

For this purpose we estimate the sum by an integral, change the variables $s = \sqrt{t}\psi(t)$ and use the fact that $t^{1/2-\delta}\psi(t)$ is increasing:

$$\begin{aligned} \sum_{j \geq n} e^{-2\sqrt{j}\psi(2j)} &\leq \int_{n/2}^\infty e^{-2\sqrt{t}\psi(2t)} dt \leq \int_n^\infty e^{-\sqrt{2t}\psi(t)} dt \\ &\leq \int_n^\infty e^{-\sqrt{2t}\psi(t)} \frac{\sqrt{t}}{\delta\psi(t)} \left(\frac{\psi(t)}{2\sqrt{t}} + \sqrt{t}\psi'(t) \right) dt \\ &\leq \int_n^\infty e^{-\sqrt{t}\psi(t)} \left(\frac{\psi(t)}{2\sqrt{t}} + \sqrt{t}\psi'(t) \right) dt \\ &= \int_{\sqrt{n}\psi(n)}^\infty e^{-s} ds = e^{-\sqrt{n}\psi(n)}. \end{aligned}$$

This proves (5.8) and gives

$$\mu^{(\alpha)}(B(w_n)) \leq_{a.a.} e^{-\sqrt{n}\psi(n)} \quad \mathbb{P}^{(\alpha)}\text{-a.s.}$$

completing the proof of the Lemma. \square

As an immediate consequence of the Lemma and the discussion above we obtain upper estimates in Theorem (2.1) and continuity of μ .

Corollary 5.3. *Under hypotheses of Theorem 2.1 the measure μ is continuous and*

$$\mu(B(w_n)) \leq_{a.a.} e^{-\sqrt{n}\psi(n)} \quad \mathbb{P}^*\text{-a.s.}$$

if $\int^\infty \psi(t)t^{-1} dt < \infty$.

5.4. A useful formula for the measure μ . In the proof of our main result the following representation of the sequence $\mu(B(w_n))$ will be needed:

Lemma 5.4. *For $\mathbb{P}^{(\alpha)}$ a.e. infinite ray $\{w_n\} \in \partial\mathbb{T}$ we have*

$$(5.9) \quad \mu(B(w_n)) = \sum_{k \geq n} e^{-V(w_k)} \hat{D}_k,$$

where

$$\hat{D}_n = c_8 \sum_{x \in C(w_n) \setminus \{w_{n+1}\}} e^{-(V(x) - V(w_n))} D_x.$$

Proof of Lemma 5.4. Because of (4.1), we have

$$\begin{aligned} \mu(B(w_n)) &= \sum_{x \in C(w_n)} \mu(B(x)) \\ &= \mu(B(w_{n+1})) + \sum_{x \in C(w_n) \setminus \{w_{n+1}\}} \mu(B(x)) \\ &= \mu(B(w_{n+1})) + c_8 e^{-V(w_n)} \sum_{x \in C(w_n) \setminus \{w_{n+1}\}} e^{-(V(x) - V(w_n))} D_x \\ &= \mu(B(w_{n+1})) + e^{-V(w_n)} \hat{D}_n. \end{aligned}$$

Notice that by iterating the formula above and $\mu(B(w_n)) \rightarrow 0$, by continuity of μ , we conclude the lemma. \square

We close this section with two more lemmas which establish that the contribution of \hat{D}_n is negligible by providing upper and lower estimates respectively.

Lemma 5.5. *Assume (2.1) and (2.3). Then for $\delta > 1/p$*

$$\hat{D}_n \leq_{a.a.} e^{n^\delta} \quad \mathbb{P}^{(\alpha)} - a.s.$$

The proof of this Lemma is exactly the same as the proof of (5.6) in Lemma 5.2. We omit details.

Lemma 5.6. *Assume (2.1). There exists $\eta > 0$ such that for all sufficiently large $\alpha \geq 0$*

$$\max_{n^3 \leq j < (n+1)^3} \hat{D}_j \geq_{a.a.} \eta \quad \mathbb{P}^{(\alpha)} - a.s.$$

Proof. By our assumptions we can infer existence of $M, \delta_0, \delta_1, \delta_2 > 0$ such that

- $\mathbb{P}[\text{there are } x \neq y \text{ such that } |x| = |y| = 1 \text{ and } V(x), V(y) \in (-M, M)] \geq \delta_0$
- $\mathbb{P}[D > \delta_1] > \delta_2$

Note that for $\alpha > 2M$, by (3.5), we have

$$\frac{R(\alpha - M)}{R(\alpha)} \geq \frac{c_1}{2c_2} =: \delta_3.$$

The following claim holds true with $\eta = \delta_1 e^{-M}$, $\delta = \delta_0 \delta_2 \delta_3 e^{-M}$

$$(5.10) \quad \mathbb{P}^{(\alpha)}[\hat{D}_1 > \eta] \geq \delta$$

for any $\alpha > 2M$. Indeed,

$$\begin{aligned}
\mathbb{P}^{(\alpha)}[\widehat{D}_1 > \eta] &= \mathbb{P}^{(\alpha)}\left[\sum_{\substack{|x|=1, \\ w_1 \neq x}} e^{-V(x)} D_x > \eta\right] \\
&\geq \mathbb{P}^{(\alpha)}\left[e^{-V(x)} D_x > \eta \text{ for some } x \neq w_1, |x| = 1\right] \\
&\geq \mathbb{P}^{(\alpha)}\left[e^{-V(x)} > e^{-M}, D_x > \delta_1 \text{ for some } x \neq w_1, |x| = 1\right] \\
&= \mathbb{E}^{(\alpha)}\left[\mathbb{P}^{(\alpha)}[V(x) < M, D_x > \delta_1 \text{ for some } x \neq w_1, |x| = 1 \mid \mathcal{F}_1]\right] \\
&\geq \mathbb{E}^{(\alpha)}\left[\mathbb{1}_{\{V(x) < M \text{ for some } x \neq w_1, |x|=1\}} \mathbb{P}[D > \delta_1]\right] \\
&\geq \delta_2 \mathbb{P}^{(\alpha)}[V(x) < M, \text{ for some } x \neq w_1, |x| = 1].
\end{aligned}$$

The remaining probability can be bounded from below by

$$\begin{aligned}
\mathbb{P}^{(\alpha)}\left[\bigcup_{\substack{|x|=1, \\ w_1 \neq x}} V(x) < M\right] &= \mathbb{E}\left[\sum_{|y|=1} \mathbb{1}_{\{\exists x \neq y V(x) < M\}} \frac{R(\alpha + V(y))}{R(\alpha)} e^{-V(y)} \mathbb{1}_{\{V(y) \geq -\alpha\}}\right] \\
&\geq \delta_3 \mathbb{E}\left[\sum_{|y|=1} \mathbb{1}_{\{\exists x \neq y V(x) < M\}} e^{-V(y)} \mathbb{1}_{\{V(y) \in (-M, M)\}}\right] \\
&\geq \delta_3 e^{-M} \mathbb{P}[\exists x \neq y V(x), V(y) \in (-M, M)] \\
&\geq \delta_0 \delta_3 e^{-M}.
\end{aligned}$$

This proves (5.10).

At the end of the proof of this lemma, we will invoke the Borel-Cantelli lemma, so first we consider the sequence

$$\begin{aligned}
\mathbb{P}^{(\alpha)}\left[\max_{k^3 \leq i < (k+1)^3} \widehat{D}_i < \eta\right] &\leq \mathbb{P}^{(\alpha)}\left[\max_{k^3 \leq i < (k+1)^3} V(w_i) < M\right] \\
&\quad + \mathbb{P}^{(\alpha)}\left[\max_{k^3 \leq i < (k+1)^3} \widehat{D}_i < \eta; \max_{k^3 \leq i < (k+1)^3} V(w_i) \geq M\right].
\end{aligned}$$

From here, we will use an induction argument in order to show that for any $k, l \in \mathbb{N}$ and $\alpha > 2M$

$$(5.11) \quad \mathbb{P}^{(\alpha)}\left[\max_{k \leq i < l} \widehat{D}_i < \eta; \max_{k \leq i < l} V(w_i) \geq M\right] \leq (1 - \delta)^{l-k+1}.$$

Put

$$A(k, l) = \left\{ \max_{k \leq i < l} \widehat{D}_i < \eta; \max_{k \leq i < l} V(w_i) \geq M \right\}.$$

For $l = k$ (5.11) follows from (5.10). For bigger l we argue that

$$\begin{aligned}
\mathbb{P}^{(\alpha)}[A(k, l+1)] &= \mathbb{E}^{(\alpha)} \left[\mathbb{P}^{(\alpha)} \left[A(k, l+1) \mid V(w_i), \widehat{D}_i, i \leq l, V(w_{l+1}) \right] \right] \\
&= \mathbb{E}^{(\alpha)} \left[\mathbb{1}_{A(k, l)} \mathbb{1}_{\{V(w_{l+1}) > M\}} \mathbb{P}^{(\alpha)} \left[\widehat{D}_{l+1} \leq \eta \mid V(w_{l+1}) \right] \right] \\
&= \mathbb{E}^{(\alpha)} \left[\mathbb{1}_{A(k, l)} \mathbb{1}_{\{V(w_{l+1}) > M\}} \mathbb{P}^{(\alpha+V(w_{l+1}))} \left[\widehat{D}_1 \leq \eta \right] \right] \\
&\leq \mathbb{P}^{(\alpha)}[A(k, l)] (1 - \delta).
\end{aligned}$$

Finally, from (5.11) and Lemma 3.4 we infer that for sufficiently large k

$$\mathbb{P}^{(\alpha)} \left[\max_{k^3 \leq j < (k+1)^3} \widehat{D}_j \leq \eta \right] \leq (1 - \delta)^{3k^2} + C \log k / k^{3/2}.$$

From this the claim follows by Borel-Cantelli lemma. \square

6. PROOFS OF THEOREMS 2.1 AND 2.2

Proof of Theorem 2.1. Step 1. Upper bound. The upper bound was already proved in Corollary 5.3

Step 2. Lower bound. In view of (5.9)

$$\mu(B(w_n)) \geq e^{-V(w_n)} \widehat{D}_n.$$

Lemma 3.2 ensures that for any ψ such that $\int^\infty \psi(t) dt/t = \infty$ we have

$$(6.1) \quad V(w_n) <_{i.o.} \frac{1}{2} \sqrt{n} \psi(n) \quad \mathbb{P}^{(\alpha)} \text{ a.s.},$$

that implies

$$(6.2) \quad e^{-V(w_n)} >_{i.o.} e^{-\frac{1}{2} \sqrt{n} \psi(n)} \quad \mathbb{P}^{(\alpha)} \text{ a.s.}$$

The idea of the proof is the following. We will prove that for some large constant M , $\widehat{D}_n > e^{-M}$, $\mathbb{P}^{(\alpha)}$ i.o., moreover this estimate holds i.o. on the set where (6.2) is satisfied. For a rigorous argument define a sequence of stopping times

$$T_0 = 0, \quad T_k = \inf \{n > T_{k-1} : 0 < V(w_n) < \sqrt{n} \psi(n)/2\}.$$

Because of (6.1), these stopping times are finite a.s. Let

$$\begin{aligned}
A_k = \{ & -\alpha < \underline{V}(x_0), -\alpha/2 < V(x_0) < M \text{ and } D_{x_0} > 1 \\
& \text{for at least one child } x_0 \text{ of } w_{T_k} \text{ different from } w_{T_{k+1}} \}.
\end{aligned}$$

Choose ε , $\alpha_0 > 0$ and $M \leq \alpha_0/2$ such that for $\alpha > \alpha_0$

$$\mathbb{P}[\text{there are } x \neq y, |x| = |y| = 1 \text{ and } -M < V(x), V(y) < M] > \varepsilon.$$

Put $\mathcal{F}_n^* = \sigma(\mathcal{F}_n, V(w_k), k \leq n)$. Then, since $A_1 \in \mathcal{F}_{T_1}^*$, for $\alpha > \alpha_0$ and $\delta = \frac{c_1}{2c_2}$

$$\begin{aligned} \mathbb{P}^{(\alpha)}[A_1] &= \mathbb{E} \left[\mathbb{1}_{A_1} \frac{D_{T_1}^{(\alpha)}}{R(\alpha)} \right] \\ &= \mathbb{E} \left[\sum_{|x|=T_1} \frac{R(\alpha + V(x))}{R(\alpha)} e^{-V(x)} \mathbb{1}_{\{V(x) > -\alpha\}} \mathbb{1}_{\{\exists y \neq x, V(y) < M\}} \right] \\ &\geq \frac{1 + \alpha - \alpha_0/2}{1 + \alpha} \varepsilon \delta e^{-M} \geq \frac{1}{2} \varepsilon \delta e^{-M}. \end{aligned}$$

Since $\mathbb{P}^{(\alpha)}[A_k \mid \mathcal{F}_{T_k}^*, V(w_{T_k}) = u] = \mathbb{P}^{(\alpha+u)}[A_1]$ we have also

$$\mathbb{P}^{(\alpha)}[A_k \mid \mathcal{F}_{T_k}^*] \geq \frac{1}{2} \varepsilon \delta e^{-M}.$$

Then by the conditioned Borel-Cantelli lemma (see e.g. Corollary 5.29 in Breiman [16]) $\mathbb{P}^{(\alpha)}[A_k \text{ i.o.}] = 1$. Hence

$$\mu(B(w_n)) \geq_{i.o.} e^{-\frac{1}{2}\sqrt{n}\psi(n)} \delta \varepsilon e^{-M} \geq_{i.o.} e^{-\sqrt{n}\psi(n)} \quad \mathbb{P}^{(\alpha)} \text{ a.s.}$$

□

Proof of Theorem 2.2. Step 1. Lower bound. As in the previous case it is sufficient to prove the result for $\mathbb{P}^{(\alpha)}$ a.e. spine $\{w_n\}$ and all sufficiently large α . Take an arbitrary $\delta > 0$. Then by Lemma 3.1

$$V(w_n) \leq_{a.a.} (1 + \delta/2) \sqrt{2\sigma^2 n \log \log n} \quad \mathbb{P}^{(\alpha)} \text{ a.s.}$$

From this we obtain

$$\mu(B(w_n)) \geq_{a.a.} \sum_{j \geq n} e^{-(1+\delta/2)\sqrt{2\sigma^2 j \log \log(j)}} \hat{D}_j.$$

Now we use Lemma 5.6 and for the sequence $(k_n)_n$ such that $(k_n - 1)^3 \leq n < k_n^3$ we write, since $\hat{D}_j \geq 0$

$$\begin{aligned} \mu(B(w_n)) &\geq_{a.a.} \left(\sum_{n \leq j < k_n^3} + \sum_{k_n^3 \leq j \leq (k_n+1)^3} + \sum_{j > (k_n+1)^3} \right) e^{-(1+\delta/2)\sqrt{2\sigma^2 j \log \log(j)}} \hat{D}_j \\ &\geq_{a.a.} \sum_{k_n^3 \leq j \leq (k_n+1)^3} e^{-(1+\delta/2)\sqrt{2\sigma^2 j \log \log(j)}} \hat{D}_j \\ &\geq_{a.a.} \eta e^{-(1+\delta/2)\sqrt{2\sigma^2 (k_n+1)^3 \log \log((k_n+1)^3)}} \\ &\geq_{a.a.} e^{-(1+\delta)^2 \sqrt{2\sigma^2 n \log \log n}}. \end{aligned}$$

This completes the first step.

Step 2. Upper bound. Fix $\delta > 0$. First we write lower estimates for the spine $\{V(w_k)\}$. Lemma 3.1 gives

$$(6.3) \quad V(w_n) >_{i.o.} (1 - \delta/8) \sqrt{2n\sigma^2 \log \log n} \quad \mathbb{P}^{(\alpha)} \text{ a.s.}$$

Choosing $\psi(n) = 1/(\log n)^2$ in Lemma 3.2 we obtain

$$(6.4) \quad V(w_n) >_{a.a.} \frac{3\sqrt{n}}{(\log n)^2} \quad \mathbb{P}^{(\alpha)} \text{ a.s.}$$

We define a sequence of stopping times (this is a subsequence of the indices for which (6.3) holds)

$$\begin{aligned} T_1 &= \inf \left\{ n : V(w_n) > (1 - \delta/8) \sqrt{2\sigma^2 n \log \log n} \right\}, \\ T_{k+1} &= \inf \left\{ n \geq T_k (\log T_k)^5 : V(w_n) > (1 - \delta/8) \sqrt{2\sigma^2 n \log \log n} \right\}. \end{aligned}$$

In view of (6.3) these stopping times are finite $\mathbb{P}^{(\alpha)}$ a.s. Denote

$$A_{k+1} = \left\{ V(w_n) \geq (1 - \delta/4) \sqrt{2\sigma^2 T_k \log \log T_k} \text{ for } T_k < n \leq T_k (\log T_k)^5 \right\}.$$

We will prove that $\mathbb{P}^{(\alpha)}[A_n \text{ i.o.}] = 1$. Notice that, applying Lemma 3.3 and (3.5) we have

$$\begin{aligned} \mathbb{P}^{(\alpha)}[A_{k+1} | \mathcal{F}_{T_k}^*] &= \mathbb{P}^{(\alpha)} \left[\min_{T_k < n \leq T_k (\log T_k)^5} V(w_n) > (1 - \delta/4) \sqrt{2\sigma^2 T_k \log \log T_k} \middle| V(w_{T_k}) \right] \\ &\geq \mathbb{P}^{(\alpha)} \left[\min_{n > T_k} V(w_n) > (1 - \delta/4) \sqrt{2\sigma^2 T_k \log \log T_k} \middle| V(w_{T_k}) \right] \\ &\geq \frac{R(\delta/8 \sqrt{2\sigma^2 T_k \log \log T_k})}{R(\alpha + (1 - \delta/8) \sqrt{2\sigma^2 T_k \log \log T_k})} > \varepsilon \delta. \end{aligned}$$

Since $A_k \in \mathcal{F}_{T_k}^*$, the conditioned Borel-Cantelli lemma (Corollary 5.29 in Breiman [16]) implies that $\mathbb{P}^{(\alpha)}[A_k \text{ i.o.}] = 1$. Therefore, by (5.8), (6.3), (6.4) and Lemma 5.5, for $n = T_k$ such that A_k holds, we have

$$\begin{aligned} \mu(B(w_n)) &\leq_{i.o.} \sum_{n \leq k \leq n(\log n)^5} e^{-(1-\delta/4)\sqrt{2\sigma^2 n \log \log n}} e^{k^\delta} + \sum_{k > n(\log n)^5} e^{-\frac{3\sqrt{k}}{(\log k)^2}} e^{k^\delta} \\ &\leq_{i.o.} n(\log n)^5 e^{-(1-\delta/4)\sqrt{2\sigma^2 n \log \log n}} e^{(n(\log n)^5)^\delta} + \sum_{k > n(\log n)^5} e^{-\frac{2\sqrt{k}}{(\log k)^2}} \\ &\leq_{i.o.} e^{-(1-\delta/2)\sqrt{2\sigma^2 n \log \log n}} + e^{-\frac{2\sqrt{n(\log n)^5}}{(\log(n(\log n)^5))^2}} \\ &\leq_{i.o.} e^{-(1-\delta)\sqrt{2\sigma^2 n \log \log n}}. \end{aligned}$$

This proves Theorem 2.2. □

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